

Bayesian nonparametric tests based on the Imprecise Dirichlet process

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Introduction

- Frequentist procedures are usually adopted in nonparametric hypothesis tests. However,
- p -value are easily misinterpreted;
- no principled way for making decisions ($\alpha=0.05$ or $\alpha=0.01$).
- Bayesian tests compute the posterior probability of the hypotheses and allow taking optimal decisions.

Need of Bayesian alternatives to frequentist nonparametric tests.

- Contradiction of Bayesian nonparametric: wish to minimize assumptions;
- require infinitesimally minute details for prior specification.

We propose a default (near-)ignorance mechanism for prior elicitation.

Dirichlet Process (DP)

- Assigns a distribution over probability distributions.
- It provides a Bayesian justification of some traditional nonparametric estimators.

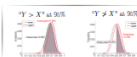
Definition:

- Let $P \sim \text{Dp}(s, g_0)$, with:
 - s = prior strength;
 - g_0 = prior base probability measure.
- Given a finite partition B_1, B_2, \dots, B_m of X , it holds $(P(B_1), \dots, P(B_m)) \sim \text{Dir}(s g_0(B_1), \dots, s g_0(B_m))$.
- $E[P(B_i)] = g_0(B_i)$
- $\text{Var}[P(B_i)] = \frac{g_0(B_i)(1-g_0(B_i))}{s+1}$

Conjugacy:

Let X_1, X_2, \dots be n samples from P

$$P|X_1, X_2, \dots \sim \text{Dp}(s+n, \frac{s}{s+n}g_0 + \frac{1}{s+n} \sum_{i=1}^n \delta_{X_i})$$



Imprecise DP (IDP)

Definition

An IDP is the set of all DP with s fixed and g_0 free to vary in the set of all probability measures \mathcal{P} .

$$\text{IDP} := \{Dp(s, g_0) : g_0 \in \mathcal{P}\}$$

$$0 = \underline{g}_0(B_i) \leq E[P(B_i)] \leq \overline{g}_0(B_i) = 1$$

No prior information about P

$$f(x) \leq f(x) \leq f(x)$$

$$f(n) - E[f|P - b_n] \leq E[f|I] \leq E[f|P - b_n] + f(n)$$

No prior information about $f(x)$

Choice of the prior strength s

- Fix the fraction of prior information after one observation OR
- Fix the minimum number of observation necessary to make a decision.

Hypothesis testing

A simple example - the sign test

$H_0: P(X < 0) \leq 0.5$, $H_1: P(X < 0) > 0.5$
 $n_0 := \#(X_i < 0)$, $n_1 := \#(X_i \geq 0)$

Lower: $\underline{P}(H_1) = \beta_0^2 \text{Beta}(n_0, s + n_0)$
Upper: $\overline{P}(H_1) = \beta_1^2 \text{Beta}(s + n_0, n_0)$

Example:
 $n_0 = 5$,
 $n_1 = 2$,
 $s = 1$



Decision making

L_0 : loss accepting H_1 (action a_1) if H_0 true
 L_1 : loss accepting H_0 (action a_0) if H_1 true

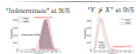
Minimum risk: accept H_1 if

$$\text{Expected loss}|a_1 < \text{Expected loss}|a_0$$
$$\Rightarrow \text{Prob}(H_1) > \frac{L_0}{L_0 + L_1} = 1 - \alpha.$$

What if $\overline{P}(H_1) > 1 - \alpha$ but

$$\underline{P}(H_1) < 1 - \alpha?$$

The decision is prior-dependent. No robust decision can be made.



Advantages

Tractability

- Simple elicitation.
- Sampling is easier than standard sampling from the DP.

Sample $P^{(k)}$ from $Dp(s+n, g_0)$

$$P^{(k)} = \text{unif}_{X_0} + \sum_{i=1}^n \text{unif}_{X_i}$$

with $(\text{un}_0, \text{un}_1, \dots, \text{un}_n) \sim \text{Dir}(s, 1, \dots, 1)$

Asymptotic consistency

The IDP tests are asymptotically consistent. Frequentist ones can have inconsistencies.

Example: Wilcoxon rank-sum test

Sample: $X_0, X_1, \dots \sim F_X; Y_1, Y_2, \dots \sim F_Y$

Hypothesis: $H_0: F_X = F_Y; H_1: F_X \neq F_Y$

Statistic: $U = \sum_{i=1}^n \sum_{j=1}^m I_{\{X_i > Y_j\}}$

- 95% rejection with $\alpha = 0.05, n = 20$, $X \sim N(0, 1)$.
- The power $\ll 1$ even for large n_1, n_2 .

- To avoid this situation, it is assumed that F_X and F_Y have the same shape.



Indecision

We empirically observed that some frequentist test can virtually behaves as a random guess when the IDP test is indeterminate.

- This shows that prior-dependent inductions are critical.
- It makes sense to suspend the decisions in these instances.



The IDP statistical package

The IDP project (in progress) has developed IDP version of

- Wilcoxon rank-sum test;
- Wilcoxon signed test;
- Sign test;
- Analysis of survival data.



<http://ppg.idisia.ch/software/IDP.php>